



Stochastic Modelling of Seasonal Influenza Dynamics: Integrating Random Perturbations and Behavioural Factors

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Abstract. The study proposes a stochastic model to investigate the seasonality of influenza in Saudi Arabia. In contrast to the classical deterministic model, we incorporate internal stochastic ambient noise through white noise perturbations in order to present a more realistic portrayal of the oscillation of sick individuals. The model correctly reproduces the empirical seasonal peak in the number of influenza cases, which is most strongly expressed in epidemiologic week 30, also showing the seasonal outbreak periodicity. Sensitivity analysis demonstrates that the magnitudes of stochastic fluctuations play a crucial role in the prediction uncertainty and the outbreak variability. Transmission rates and the recovery parameter are the main determinants of the magnitude, timing, and impact on hospitalization of the epidemic wave. A greater transmission rate is consistently associated with more intense and prolonged breakouts, whereas a larger transmission rate facilitates a more rapid ascent and descent of the peak. These findings underscore the significance of stochastic factors and pinpoint essential parameters for enhancing public health interventions and resource allocation strategies. Our results provide practical recommendations for enhancing influenza preparation using data-driven stochastic modelling.

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1. Introduction

Seasonal flu is an infectious viral respiratory disease that spreads during the autumn and winter seasons [1]. The illness comes with typical symptoms such as fever, headache, chills, cough, nasal congestion, body aches and pains. It has been demonstrated that in severe cases of seasonal influenza, mortality can occur [1].

Persons with multiple chronic medical conditions are more likely to become infected with seasonal influenza [2] and experience an increase in its severity with higher rates of emergency room visits and hospitalizations [3]. Seasonal influenza is responsible for killing between 290,000 and 650,000 people annually, according to the WHO in 2022 [4].

In 2020, World Health Ranking estimated that influenza and pneumonia, together, caused the death of 4.58% of the population of Saudi Arabia and its seasonal influenza death rate was 30.9 per 100 000, so that it is 82nd most-frequented country in the world [5].

A systematic review and meta-analysis of 18 studies in Hajj pilgrims reporting on 62,431 individuals estimated a prevalence of 5.9% for influenza A and 3.6% for influenza B [6].

Among currently approved viral vaccines, vaccine for the influenza virus is distinctive. Vaccines that target influenza disease have to be reformulated annually to provide matching between the vaccine strain and the wild-type viruses that will be present in a season and this is owing to the influenza virus's ongoing antigenic drift of its HA and NA surface glycoproteins. Licensed inactivated influenza vaccines (IIV) have been available in the United States (US) since 1945 [7]. A live attenuated influenza vaccine has been available since it was approved in 2003 [8]. The objective of the WHO Global Vaccine Action Plan (GVAP) is to reduce transmission of this global epidemic by increasing the seasonal influenza vaccine coverage rate [9].

Annual vaccination against influenza is the most effective measure to prevent infections and to reduce the adverse events associated with them, such as visit to the doctor, hospitalization, and deaths [10]. There is variation in VE from year to year depending on the degree of match between the vaccine viruses and those in circulation. Accordingly, vaccine effectiveness (VE) could be weakened by lack of compatibility with the strains of influenza which are presently spreading and, hence, E_t of the influenza-related outcomes increases [11–14].

However studies suggest that the uptake of seasonal influenza vaccine in Saudi Arabia is low. The references included six Saudi citizens' studies published from October 19, 2017, to October 18, 2022. Seasonal influenza vaccine coverage was 12.7%-55.0% in these studies [15–20]. Taken together, these studies exemplify that seasonal flu vaccination coverage is below that recommended by vaccination guidelines, despite the different methods of measuring vaccine status. As such, monitoring the status of seasonal influenza vaccination should be regularly implemented among the residents of the Saudi population in order to guide the efforts of boosting vaccination coverage.

Due to the increasing use of fractional calculus in modeling diffusion, control processes, and viscoelasticity, applied mathematics has gained notable popularity over the

past few decades, particularly in physics and engineering research, where fractional differential equations are widely employed [21–25]. This mathematical discovery has had considerable impact on the practice of epidemiological modelling, given that models play a key role in evaluating the transmission dynamics of infectious diseases [26–30]. The majority of current models are classical deterministic and fractal models, which may ignore the inherent natural variability of disease transmission and the background noise or ambient noise. Environmental randomness is an important factor in the occurrence and progression of epidemics, and there is a greater attention to stochastic models containing randomness.

Stochastic systems and fractional-order mathematical models have been more widely used in the study of the transmission dynamics of infectious diseases in recent years. For example, fractional differential equations and stochastic systems have been used in modeling complex diseases such as visceral leishmaniasis, flu pandemics, and COVID-19. These methodologies enable better understanding of disease persistence, extinction conditions, and evaluation of control strategies [31–37].

Ali, M. et al. [07] proposed a new stochastic model change point SEIHR to predict COVID-19 in South Africa and to evaluate the effectiveness of NPIs. The study developed solutions using stochastic Lyapunov function theory and analyzed data gathered between April and September 2021 to explore the dynamics of the pandemic. The model tests the efficacy of lockdowns in preventing the spread of viruses.

Alnafisah and El-Shahed [08] proposed a stochastic Hantavirus infection model. It was established for the solution in the space determined space. They also established conditions for a unique ergodic stationary distribution and conditions for extinction of Hantavirus infection. With the Milstein method, they highlighted the role of environmental noise in the model.

Liu and Jiang [09] considered a SIR epidemic model with logistic birth rate and possibilistic formulation, and established the global stability of the positive equilibrium by using the analysing method based on the Lyapunov functions and constructed a globally asymptotically stable solution. The stochastic basic reproduction number R_S is shown to play a role in determining the threshold dynamics of the system in given circumstances, and sufficient conditions are established for the extinction of the disease and the stable existence of the positive solutions.

Kang et al. [38] studied the SIS model in randomly fluctuating environments. They used parameter perturbation to investigate the impact of uncertainties in parameters, such as infection rate, recovery rate. They demonstrated that Stratonovich's SDE is not well defined once the fluctuation is Gaussian. The reason is that the Stratonovich SDE is more appropriate for the parameter changes in the epidemic model. In contrast, the influence of the Itô SDE scales with the variance.

We take one step further in influenza transmission modeling by incorporating stochasticity in contrast to the classical deterministic modeling. The study investigates what the intensity of white noise does to variability, similar to what happens with social behavior and environmental change. Looking at different levels of noise, the model includes the randomness of the trends in the epidemics and finds the basic parameters (

beta,
alpha) that determine the timing, severity, and maximum hospitalization of the outbreaks. The results confirm that stochastic modeling is essential for a successful public health intervention.

2. Model formulation

We consider an SEIHVR epidemic model, where the total human population at any time t is denoted by $N(t)$. subdivides into six compartments Susceptible individuals (S), exposed (E), infected (I), hospitalized (H), vaccinated (V), and recovered (R). Thus,

$$N(t) = S(t) + E(t) + I(t) + H(t) + V(t) + R(t).$$

Each class incurs a fixed natural death at rate μ . The susceptible class is assumed to be increasing by recruitment process at rate λ and by waning immunity of the individuals in classes V and R at rate ρ . It is decreased by vaccination at rate α and by the force of infection

$$f = \beta \frac{(E + I)(S + \sigma V)}{N},$$

where σ is vaccination inefficacy, i.e., $1 - \sigma$ is vaccine efficacy, and β is effective contact rate of a susceptible or vaccinated individual with an infected or exposed individual.

Individuals in the exposed class are recruited by the force of infection f , and move to class I (after spending the incubation period) at rate w and to class R at rate γ_1 . Vaccination does not confer a long lasting immunity nor full protection is guaranteed. Individuals in class V are recruited from class S at rate α , exposed to the virus at rate σ , and become susceptible again at rate ρ due to waning immunity. Infected individuals are increased at rate ω from class E , recover at rate γ , and ϵ fraction of them are hospitalized. The recovered population is increased from classes E, I , and H at rates γ_1, γ_2 , and γ_3 , respectively; and become susceptible again at rate ρ .

From the above description we derive system (1) of ordinary differential equations to mathematically describe the transmission dynamics of influenza disease.

$$\begin{aligned} \frac{dS}{dt} &= \lambda + \rho(V + R) - \beta \frac{(E + I)(S + \sigma V)}{N} - (\alpha + \mu)S \\ \frac{dE}{dt} &= \beta \frac{(E + I)(S + \sigma V)}{N} - (\omega + \gamma_1 + \mu)E \\ \frac{dI}{dt} &= \omega E - (\gamma_2 + \epsilon + \mu)I \\ \frac{dH}{dt} &= \epsilon I - (\gamma_3 + \mu)H \\ \frac{dV}{dt} &= \alpha S - (\rho + \sigma + \mu)V \\ \frac{dR}{dt} &= \gamma_1 E + \gamma_2 I + \gamma_3 H - (\rho + \mu)R \end{aligned} \tag{1}$$

with initial conditions

$$S(0) \geq 0, E(0) \geq 0, I(0) \geq 0, H(0) \geq 0, V(0) \geq 0, R(0) \geq 0.$$

Table 1: Description Parameters in the model

Parameter	Description
λ	Recruitment rate of individuals into the population
μ	Natural death rate
β	Average effective contact rate
α	Vaccination rate
σ	Vaccine inefficacy
ϵ	Fraction of infected individuals who hospitalized
γ_1	Recovery rate for exposed individuals
γ_2	Recovery rate for infected individuals
γ_3	Recovery rate for hospitalized individuals
$\frac{1}{\omega}$	Average latent or incubation period
ρ	Rate at which individuals lose immunity

The stochastic differential equation model (2) correspondence of the deterministic model (1) is given by

$$\begin{aligned}
 dS(t) &= \left(\lambda + \rho(V + R) - \beta \frac{(E + I)(S + \sigma V)}{N} - (\alpha + \mu)S \right) dt \\
 &\quad + \delta_1 S(t) dB_1(t) \\
 dE(t) &= \left(\beta \frac{(E + I)(S + \sigma V)}{N} - (\omega + \gamma_1 + \mu)E \right) dt + \delta_2 E(t) dB_2(t) \\
 dI(t) &= (\omega E - (\gamma_2 + \epsilon + \mu)I) dt + \delta_3 I(t) dB_3(t) \\
 dH(t) &= (\epsilon I - (\gamma_3 + \mu)H) dt + \delta_4 H(t) dB_4(t) \\
 dV(t) &= (\alpha S - (\rho + \sigma + \mu)V) dt + \delta_5 V(t) dB_5(t) \\
 dR(t) &= (\gamma_1 E + \gamma_2 I + \gamma_3 H - (\rho + \mu)R) dt + \delta_6 R(t) dB_6(t)
 \end{aligned} \tag{2}$$

Where $B_1(t), B_2(t), B_3(t), B_4(t), B_5(t), B_6(t)$ as independent standard Brownian motions, and $\delta_1, \delta_2, \delta_3, \delta_4, \delta_5, \delta_6$ as the intensities of the standard Gaussian white noises, respectively.

3. Analysis of the Model

3.1. Existence and uniqueness of solution to the stochastic model

This section discusses the existence and uniqueness of solution of the proposed stochastic model (2).

Lemma 1. ([39]) *For any initial condition $(S(0), E(0), I(0), H(0), V(0), R(0)) \in \mathbb{R}_+^6$, there exists a unique solution $(S(t), E(t), I(t), H(t), V(t), R(t))$ of stochastic model (2) on $t \geq 0$, which remains in \mathbb{R}_+^6 with probability one for all $t \geq 0$.*

Proof. As for initial condition of the state variables $(S(0), E(0), I(0), H(0), V(0), R(0)) \in \mathbb{R}_+^6$, the coefficients used in equations are continuous and local Lipschitz condition. Hence, there must exists a local unique solution $(S(t), E(t), I(t), H(t), V(t), R(t))$ of the stochastic model (2) over $t \in [0, \tau_e)$, τ_e is the explosion time. In order to prove that the solution is global, we only need to prove that $\tau_e = \infty$ a.s. Assume that we suppose that there exists z_0 such that all of the initial conditions on the state lie within $\{\frac{1}{z_0}, z_0\}$. Let us define for each positive integer $z \geq z_0$, define the finishing time as follows

$$\tau_z = \left\{ t \in [0, \tau_e) : \min\{S, E, I, H, V, R\} \leq \frac{1}{z} \text{ or } \max\{S, E, I, H, V, R\} \geq z \right\}$$

Let $\inf \phi = \infty$ where ϕ denotes the null set. Definition of τ_z and as $z \rightarrow \infty$, we say that it is increasing.

Assume $\tau_\infty = \lim_{z \rightarrow \infty} \tau_z$, obviously $\tau_\infty \leq \tau_e$ a.s. Upon showing $\tau_\infty = \infty$ a.s., we suppose that $\tau_e = \infty$ and hence $(S(t), E(t), I(t), H(t), V(t), R(t))$ will lie in \mathbb{R}_+^6 a.s. $\forall t \geq 0$. Hence, it is sufficient to prove that $\tau_e = \infty$ a.s. if not, there must exists two positive constants $\zeta \in (0, 1)$ and T such that $P\{T \geq \tau_\infty\} > \zeta$, we can find an integer $z_1 \geq z_0$, such that $P\{\tau_z \leq T\} \geq \zeta, \forall z \geq z_1$. Define a \mathbb{C}^2 -function $G: \mathbb{R}_+^6 \rightarrow \mathbb{R}_+$ by

$$G(S(t), E(t), I(t), H(t), V(t), R(t)) = S + E + I + H + V + R - \ln S - 6 - \ln E - \ln I - \ln H - \ln V - \ln R \quad (3)$$

We note that G is a non-negative function, such that $0 \leq g - \ln g - 1, \forall g > 0$. Assume that $z_0 \leq z$ and $T > 0$ are arbitrary. Upon applying Itô's formula to equation (3), So we get

$$\begin{aligned} dG(S, E, I, H, V, R) &= LG(S, E, I, H, V, R) + \delta_1(S - 1)dB_1 + \delta_2(E - 1)dB_2 \\ &\quad + \delta_3(I - 1)dB_3 + \delta_4(H - 1)dB_4 + \delta_5(V - 1)dB_5 \\ &\quad + \delta_6(R - 1)dB_6 \end{aligned} \quad (4)$$

In equation (4), $LG(S, E, I, H, V, R): \mathbb{R}_+^6 \rightarrow \mathbb{R}_+$ is defined by

$$\begin{aligned} LG &= \left(1 - \frac{1}{S}\right) \left(\lambda + \rho(V + R) - \beta \frac{(E + I)(S + \sigma V)}{N} - (\alpha + \mu)S \right) \\ &\quad + \left(1 - \frac{1}{E}\right) \left(\beta \frac{(E + I)(S + \sigma V)}{N} - (\omega + \gamma_1 + \mu)E \right) \\ &\quad + \left(1 - \frac{1}{I}\right) (\omega E - (\gamma_2 + \epsilon + \mu)I) \\ &\quad + \left(1 - \frac{1}{H}\right) (\epsilon I - (\gamma_3 + \mu)H) \\ &\quad + \left(1 - \frac{1}{V}\right) (\alpha S - (\rho + \sigma + \mu)V) \\ &\quad + \left(1 - \frac{1}{R}\right) (\gamma_1 E + \gamma_2 I + \gamma_3 H - (\rho + \mu)R) \end{aligned}$$

$$\begin{aligned}
& + \frac{\delta_1^2 + \delta_2^2 + \delta_3^2 + \delta_4^2 + \delta_5^2 + \delta_6^2}{2} \\
\leq & \lambda + \alpha + 5\mu + \omega + \gamma_1 + \gamma_2 + \gamma_3 + \epsilon + \sigma + 2\rho + \beta + \beta\sigma \\
& + \frac{\delta_1^2 + \delta_2^2 + \delta_3^2 + \delta_4^2 + \delta_5^2 + \delta_6^2}{2} = K
\end{aligned} \tag{5}$$

Then

$$\begin{aligned}
dG(X) \leq & Kdt + \delta_1(S-1)dB_1 + \delta_2(E-1)dB_2 + \delta_3(I-1)dB_3 \\
& + \delta_4(H-1)dB_4 + \delta_5(V-1)dB_5 + \delta_6(R-1)dB_6
\end{aligned}$$

Integrating both sides from 0 to $\tau_z \wedge T$, and taking expectations, we obtain

$$\begin{aligned}
& E[G(S(\tau_z \wedge T)), E(\tau_z \wedge T), I(\tau_z \wedge T), H(\tau_z \wedge T), V(\tau_z \wedge T), R(\tau_z \wedge T)] \\
& \leq G(S(0), E(0), I(0), H(0), V(0), R(0)) + E\left(\int_0^{\tau_z \wedge T} Kdt\right) \\
& \leq G(S(0), E(0), I(0), H(0), V(0), R(0)) + KT
\end{aligned} \tag{6}$$

Then

$$E[G] \leq G(S(0), E(0), I(0), H(0), V(0), R(0)) + KT$$

for any positive $z \geq z_1$, we set $\Omega_z = (\tau_z < T)$, this leads to $P(\Omega_z) \geq \zeta$.

Note that for each $\omega \in \Omega_z$ there must exist one more than one $S(\tau_z, \omega), E(\tau_z, \omega), I(\tau_z, \omega), H(\tau_z, \omega), V(\tau_z, \omega), R(\tau_z, \omega)$ which equals $\frac{1}{z}$ or z , Consequently, $G(S(\tau_z), E(\tau_z), I(\tau_z), H(\tau_z), V(\tau_z), R(\tau_z))$ is no less than $\frac{1}{z} - 1 + \ln z$ or $z - 1 - \ln z$. Therefore,

$$G(S(\tau_z), E(\tau_z), I(\tau_z), H(\tau_z), V(\tau_z), R(\tau_z)) \geq \left(\frac{1}{z} - 1 + \ln z\right) \wedge (z - 1 - \ln z)$$

So we obtain

$$\begin{aligned}
G(S(0), E(0), I(0), H(0), V(0), R(0)) + KT & \geq E[I_{\Omega_\omega} G(S(\tau_z), E(\tau_z), \\
& I(\tau_z), H(\tau_z), V(\tau_z), R(\tau_z))] = P(\Omega_K) G(S(\tau_z), E(\tau_z), I(\tau_z), \\
& H(\tau_z), V(\tau_z), R(\tau_z)) > \zeta \left[\left(\frac{1}{z} - 1 + \ln z\right) \wedge (z - 1 - \ln z)\right]
\end{aligned}$$

I_{Ω_ω} is the indicator function of Ω_ω . Set $z \rightarrow \infty$, we have

$$\infty > G(S(0), E(0), I(0), H(0), V(0), R(0)) + KT > \infty$$

showing that $t_\infty = \infty, a.s.$

Lemma 2. For any positive solution $(S(t), E(t), I(t), H(t), V(t), R(t))$ of stochastic model (2) with initial value $(S(0), E(0), I(0), H(0), V(0), R(0)) \in \mathbb{R}_+^6$ we have

$$\begin{aligned}
& \max\{\lim_{t \rightarrow \infty} \sup S(t), \lim_{t \rightarrow \infty} \sup E(t), \lim_{t \rightarrow \infty} \sup I(t), \lim_{t \rightarrow \infty} \sup H(t), \\
& \lim_{t \rightarrow \infty} \sup V(t), \lim_{t \rightarrow \infty} \sup R(t)\} \leq \frac{\lambda}{\mu}, a.s.
\end{aligned}$$

Proof. From stochastic model (2), we have

$$\begin{aligned} \lim_{t \rightarrow \infty} \frac{d(S(t) + E(t) + I(t) + H(t) + V(t) + R(t))}{dt} &= \lim_{t \rightarrow \infty} (\Lambda - \mu(S(t) \\ &\quad + E(t) + I(t) + H(t) \\ &\quad + V(t) + R(t)) - \sigma V) \\ \lim_{t \rightarrow \infty} (S(t) + E(t) + I(t) + H(t) + V(t) + R(t)) &\leq \frac{\lambda}{\mu} - \frac{\lambda}{\mu} e^{-\mu t} \leq \frac{\lambda}{\mu} \end{aligned}$$

Then obviously we obtain

$$\begin{aligned} \lim_{t \rightarrow \infty} \sup S(t) &\leq \frac{\lambda}{\mu}, \lim_{t \rightarrow \infty} \sup E(t) \leq \frac{\lambda}{\mu}, \lim_{t \rightarrow \infty} \sup I(t) \leq \frac{\lambda}{\mu} \\ \lim_{t \rightarrow \infty} \sup H(t) &\leq \frac{\lambda}{\mu}, \lim_{t \rightarrow \infty} \sup V(t) \leq \frac{\lambda}{\mu}, \lim_{t \rightarrow \infty} \sup R(t) \leq \frac{\lambda}{\mu}, a.s. \end{aligned}$$

3.2. Extinction

In this section, we investigate the conditions for the extinction of the disease in the stochastic model (2) under the white noise stochastic disturbance. The following theorem gives conditions for $I(t)$ and $H(t)$ got extinction.

Theorem 1. *If $(\gamma_2 + \epsilon + \mu) + \frac{1}{2}\delta_3^2 > \frac{\omega\lambda}{\mu}$, then $I(t)$ go to extinction almost surely. If $(\gamma_3 + \mu) + \frac{1}{2}\delta_4^2 > \frac{\epsilon\lambda}{\mu}$, then $H(t)$ go to extinction almost surely.*

Proof. Let $(S(t), E(t), I(t), H(t), V(t), R(t))$ be a solution of stochastic model (2) with initial value $(S(0), E(0), I(0), H(0), V(0), R(0)) \in \mathbb{R}_+^6$. Applying Itô's formula to the third equation of stochastic model (2), we get

$$\begin{aligned} dI(t) &= \left(\frac{\omega E}{I} - (\gamma_2 + \epsilon + \mu) \right) I(t) dt + \delta_3 I(t) dB_3(t) \\ d(\ln I(t)) &= \left(f(t) - \frac{1}{2}g^2(t) \right) dt + g(t)dB(t) \\ &= \left(\frac{\omega E}{I} - (\gamma_2 + \epsilon + \mu) - \frac{1}{2}\delta_3^2 \right) dt + \delta_3 dB_3(t) \end{aligned}$$

integration from 0 to t , then

$$\begin{aligned} \ln I(t) - \ln I(0) &= \int_0^t \left(\frac{\omega E}{I} - (\gamma_2 + \epsilon + \mu) - \frac{1}{2}\delta_3^2 \right) du + \int_0^t \delta_3 dB_3 \\ \ln I(t) &\leq \ln I(0) + \int_0^t \left(\frac{\omega\lambda}{\mu} - (\gamma_2 + \epsilon + \mu) - \frac{1}{2}\delta_3^2 \right) du + \int_0^t \delta_3 dB_3 \\ &\leq \ln I(0) - \left((\gamma_2 + \epsilon + \mu) + \frac{1}{2}\delta_3^2 - \frac{\omega\lambda}{\mu} \right) t + \delta_3 B_3 \end{aligned}$$

Dividing by t , we have

$$\frac{\ln I(t)}{t} \leq \frac{\ln I(0)}{t} - \left((\gamma_2 + \epsilon + \mu) + \frac{1}{2}\delta_3^2 - \frac{\omega\lambda}{\mu} \right) + \frac{\delta_3 B_3}{t}$$

By using the strong law of large numbers [40], we obtain

$$\lim_{t \rightarrow \infty} \frac{\delta_3 B_3}{t} = 0$$

$$\lim_{t \rightarrow \infty} \frac{\ln I(t)}{t} \leq - \left((\gamma_2 + \epsilon + \mu) + \frac{1}{2}\delta_3^2 - \frac{\omega\lambda}{\mu} \right) \quad (7)$$

Since $(\gamma_2 + \epsilon + \mu) + \frac{1}{2}\delta_3^2 > \frac{\omega\lambda}{\mu}$, taking the limit superior of both sides leads to

$$\limsup_{t \rightarrow \infty} \frac{\ln I(t)}{t} \leq - \left((\gamma_2 + \epsilon + \mu) + \frac{1}{2}\delta_3^2 - \frac{\omega\lambda}{\mu} \right) < 0, a.s. \quad (8)$$

which implies $\lim_{t \rightarrow \infty} I(t) = 0$

From the fourth equation of system

$$\begin{aligned} dH(t) &= \left(\frac{\epsilon I}{H} - (\gamma_3 + \mu) \right) H(t)dt + \delta_4 H(t)dB_4(t) \\ d(\ln H(t)) &= \left(\frac{\epsilon I}{H} - (\gamma_3 + \mu) - \frac{1}{2}\delta_4^2 \right) dt + \delta_4 dB_4(t) \end{aligned}$$

integration from 0 to t , then

$$\begin{aligned} \ln H(t) - \ln H(0) &= \int_0^t \left(\frac{\epsilon I}{H} - (\gamma_3 + \mu) - \frac{1}{2}\delta_4^2 \right) du + \int_0^t \delta_4 dB_4 \\ \ln H(t) &\leq \ln H(0) - \left((\gamma_3 + \mu) + \frac{1}{2}\delta_4^2 - \frac{\epsilon\lambda}{\mu} \right) t + \delta_4 B_4 \end{aligned}$$

Dividing by t , we have

$$\begin{aligned} \frac{\ln H(t)}{t} &\leq \frac{\ln H(0)}{t} - \left((\gamma_3 + \mu) + \frac{1}{2}\delta_4^2 - \frac{\epsilon\lambda}{\mu} \right) + \frac{\delta_4 B_4}{t} \\ \lim_{t \rightarrow \infty} \frac{\ln H(t)}{t} &\leq - \left((\gamma_3 + \mu) + \frac{1}{2}\delta_4^2 - \frac{\epsilon\lambda}{\mu} \right) \quad (9) \end{aligned}$$

Since $(\gamma_3 + \mu) + \frac{1}{2}\delta_4^2 > \frac{\epsilon\lambda}{\mu}$, taking the limit superior of both sides leads to

$$\limsup_{t \rightarrow \infty} \frac{\ln H(t)}{t} \leq - \left((\gamma_3 + \mu) + \frac{1}{2}\delta_4^2 - \frac{\epsilon\lambda}{\mu} \right) < 0, a.s. \quad (10)$$

which implies $\lim_{t \rightarrow \infty} H(t) = 0$

3.3. Persistence and Stationary Distributions

As for as the stochastic model are concerned, they have no endemic equilibrium. Thus, the stability analysis cannot be used as a tool for studying the disease persistence. As a result, we shall establish sufficient conditions for the existence of a unique ergodic stationary distribution which in some sense, will work for persistence of the disease.

Define a parameter

$$\mathcal{R}_0^* = \frac{\sigma\beta\omega\alpha}{(\alpha + \mu + \frac{1}{2}\delta_1^2)(\omega + \gamma_1 + \mu + \frac{1}{2}\delta_2^2)(\gamma_2 + \epsilon + \mu + \frac{1}{2}\delta_3^2)(\rho + \sigma + \mu + \frac{1}{2}\delta_5^2)}$$

Theorem 2. *The solution $(S(t), E(t), I(t), H(t), V(t), R(t))$ of the stochastic model (2) is ergodic as well as there is a unique stationary distribution $\pi(\cdot)$ whenever $\mathcal{R}_0^* > 1$.*

Proof. In view of Lemma (1) we have obtained that for any initial value $(S(0), E(0), I(0), H(0), V(0), R(0)) \in \mathbb{R}_+^6$, there is a unique solution $(S(t), E(t), I(t), H(t), V(t), R(t)) \in \mathbb{R}_+^6$. The diffusion matrix of stochastic model (2) is given by

$$A = \begin{pmatrix} \delta_1^2 S^2 & 0 & 0 & 0 & 0 & 0 \\ 0 & \delta_2^2 E^2 & 0 & 0 & 0 & 0 \\ 0 & 0 & \delta_3^2 I^2 & 0 & 0 & 0 \\ 0 & 0 & 0 & \delta_4^2 H^2 & 0 & 0 \\ 0 & 0 & 0 & 0 & \delta_5^2 V^2 & 0 \\ 0 & 0 & 0 & 0 & 0 & \delta_6^2 R^2 \end{pmatrix} \quad (11)$$

If we choose

$$M = \min_{(S, E, I, H, V, R) \in \bar{D} \subset \mathbb{R}_+^6} \{ \delta_1^2 S^2, \delta_2^2 E^2, \delta_3^2 I^2, \delta_4^2 H^2, \delta_5^2 V^2, \delta_6^2 R^2 \}$$

We obtain

$$\begin{aligned} \sum_{i,j=0}^6 a_{ij}(S, E, I, H, V, R) \xi_i \xi_j &\geq \delta_1^2 S^2 \xi_1^2 + \delta_2^2 E^2 \xi_2^2 + \delta_3^2 I^2 \xi_3^2 \\ &\quad + \delta_4^2 H^2 \xi_4^2 + \delta_5^2 V^2 \xi_5^2 + \delta_6^2 R^2 \xi_6^2 \\ &\geq M |\xi|^2 \end{aligned}$$

where $(S, E, I, H, V, R) \in \bar{D}$, $\xi = \{\xi_1, \xi_2, \xi_3, \xi_4, \xi_5, \xi_6\} \in \mathbb{R}_+^6$. Then the first condition in Lemma 4.1 in [41] is satisfied.

To prove the second condition we construct a \mathbb{C}^2 -function $\mathbb{V}: \mathbb{R}_+^6 \rightarrow \mathbb{R}_+$ in the following form

$$\begin{aligned} \mathbb{V}_1(S, E, I, H, V, R) &= S + E + I + H + V + R - w_1 \ln S \\ &\quad - w_2 \ln E - w_3 \ln I - w_4 \ln V \end{aligned} \quad (12)$$

where w_1, w_2, w_3, w_4 are the positive constant . Using Itô's formula, we have

$$\begin{aligned}
 L\mathbb{V}_1 &= w_1 \left(-\frac{\lambda}{S} - \frac{\rho(V+R)}{S} + \beta \frac{(E+I)(S+\sigma V)}{SN} + (\alpha + \mu) + \frac{1}{2}\delta_1^2 \right) \\
 &\quad + \lambda - \mu N - \sigma V + w_2 \left(-\beta \frac{(E+I)(S+\sigma V)}{EN} + (\omega + \gamma_1 + \mu) + \frac{1}{2}\delta_2^2 \right) \\
 &\quad + w_3 \left(-\frac{\omega E}{I} + (\gamma_2 + \epsilon + \mu) + \frac{1}{2}\delta_3^2 \right) + w_4 \left(-\frac{\alpha S}{V} + (\rho + \sigma + \mu) + \frac{1}{2}\delta_5^2 \right) \\
 &\leq \lambda - 5 \left(\sigma V \times \frac{w_1 \lambda}{S} \times \frac{w_2 \beta I}{E} \times \frac{w_3 \omega E}{I} \times \frac{w_4 \alpha S}{V} \right)^{\frac{1}{5}} - w_2 \beta \\
 &\quad + w_1 \left(-\frac{\rho(V+R)}{S} + \beta \frac{(E+I)(S+\sigma V)}{SN} + (\alpha + \mu) + \frac{1}{2}\delta_1^2 \right) \\
 &\quad + w_2 \left((\omega + \gamma_1 + \mu) + \frac{1}{2}\delta_2^2 \right) + w_3 \left((\gamma_2 + \epsilon + \mu) + \frac{1}{2}\delta_3^2 \right) \\
 &\quad + w_4 \left((\rho + \sigma + \mu) + \frac{1}{2}\delta_5^2 \right)
 \end{aligned}$$

Using the inequality $x + y \geq 2\sqrt{xy}$, $x, y > 0$, leads to

$$\begin{aligned}
 L\mathbb{V}_1 &\leq -5 (\sigma w_1 \lambda w_2 \beta w_3 \omega w_4 \alpha)^{\frac{1}{5}} + \lambda + \frac{w_1 \beta (E+I)(S+\sigma V)}{SN} \\
 &\quad + w_1 \left(\alpha + \mu + \frac{1}{2}\delta_1^2 \right) + w_2 \left(\omega + \gamma_1 + \mu + \frac{1}{2}\delta_2^2 \right) \\
 &\quad + w_3 \left(\gamma_2 + \epsilon + \mu + \frac{1}{2}\delta_3^2 \right) + w_4 \left(\rho + \sigma + \mu + \frac{1}{2}\delta_5^2 \right)
 \end{aligned}$$

let

$$\begin{aligned}
 w_1 &= \frac{\lambda}{(\alpha + \mu + \frac{1}{2}\delta_1^2)}, w_2 = \frac{\lambda}{(\omega + \gamma_1 + \mu + \frac{1}{2}\delta_2^2)} \\
 w_3 &= \frac{\lambda}{(\gamma_2 + \epsilon + \mu + \frac{1}{2}\delta_3^2)}, w_4 = \frac{\lambda}{(\rho + \sigma + \mu + \frac{1}{2}\delta_5^2)}
 \end{aligned}$$

Then

$$\begin{aligned}
 L\mathbb{V}_1 &\leq \frac{w_1 \beta (E+I)(S+\sigma V)}{SN} - 5\lambda \left(\left(\frac{\sigma \beta}{(\alpha + \mu + \frac{1}{2}\delta_1^2)(\omega + \gamma_1 + \mu + \frac{1}{2}\delta_2^2)} \right. \right. \\
 &\quad \left. \left. \times \frac{\omega \alpha}{(\gamma_2 + \epsilon + \mu + \frac{1}{2}\delta_3^2)(\rho + \sigma + \mu + \frac{1}{2}\delta_5^2)} \right)^{\frac{1}{5}} - 1 \right)
 \end{aligned}$$

let

$$\mathcal{R}_0^* = \frac{\sigma \beta \omega \alpha}{(\alpha + \mu + \frac{1}{2}\delta_1^2)(\omega + \gamma_1 + \mu + \frac{1}{2}\delta_2^2)(\gamma_2 + \epsilon + \mu + \frac{1}{2}\delta_3^2)(\rho + \sigma + \mu + \frac{1}{2}\delta_5^2)}$$

Then

$$L\mathbb{V}_1 \leq -5\lambda(\mathcal{R}_0^{*\frac{1}{5}} - 1) + \frac{w_1 \beta (E+I)(S+\sigma V)}{SN} \quad (13)$$

In addition, we obtain

$$\begin{aligned}
 \mathbb{V}_2(S, E, I, H, V, R) &= w_5(S + E + I + H + V + R - w_1 \ln S - w_2 \ln E \\
 &\quad - w_3 \ln I - w_4 \ln V) - \ln S - \ln H - \ln R + S + E \\
 &\quad + I + H + V + R \\
 &= (w_5 + 1)(S + E + I + H + V + R) \\
 &\quad - (w_1 w_5 + 1) \ln S - w_2 w_5 \ln E - w_3 w_5 \ln I \\
 &\quad - w_4 w_5 \ln V - \ln H - \ln R
 \end{aligned}$$

Where w_5 is positive constant. It is easy to check that Notice that, let $z \rightarrow \infty$, then we have

$$\lim_{z \rightarrow \infty, (S, E, I, H, V, R) \in \mathbb{R}_+^6 \setminus U_z} \inf \mathbb{V}_2(S, E, I, H, V, R) = +\infty \quad (14)$$

where $U_z = (\frac{1}{z}, z) \times (\frac{1}{z}, z) \times (\frac{1}{z}, z) \times (\frac{1}{z}, z) \times (\frac{1}{z}, z) \times (\frac{1}{z}, z)$. The next step is to prove that $\mathbb{V}_2(S, E, I, H, V, R)$ has one and only one minimum value $\mathbb{V}_2(S(0), E(0), I(0), H(0), V(0), R(0))$. The partial derivative of $\mathbb{V}_2(S, E, I, H, V, R)$ with respect to S, E, I, H, V, R is as follow

$$\begin{aligned}
 \frac{\partial \mathbb{V}_2(S, E, I, H, V, R)}{\partial S} &= 1 + w_5 - \frac{(w_5 w_1 + 1)}{S} \\
 \frac{\partial \mathbb{V}_2(S, E, I, H, V, R)}{\partial E} &= 1 + w_5 - \frac{w_5 w_2}{E} \\
 \frac{\partial \mathbb{V}_2(S, E, I, H, V, R)}{\partial I} &= 1 + w_5 - \frac{w_5 w_3}{I} \\
 \frac{\partial \mathbb{V}_2(S, E, I, H, V, R)}{\partial H} &= 1 + w_5 - \frac{1}{H} \\
 \frac{\partial \mathbb{V}_2(S, E, I, H, V, R)}{\partial V} &= 1 + w_5 - \frac{w_5 w_4}{V} \\
 \frac{\partial \mathbb{V}_2(S, E, I, H, V, R)}{\partial R} &= 1 + w_5 - \frac{1}{R}
 \end{aligned}$$

We can easily show that \mathbb{V}_2 has unique stagnation point

$$\begin{aligned}
 (S(0), E(0), I(0), H(0), V(0), R(0)) &= \\
 \left(\frac{w_5 w_1 + 1}{1 + w_5}, \frac{w_5 w_2}{1 + w_5}, \frac{w_5 w_3}{1 + w_5}, \frac{1}{1 + w_5}, \frac{w_5 w_4}{1 + w_5}, \frac{1}{1 + w_5} \right) & \quad (15)
 \end{aligned}$$

Moreover, the Hessian matrix of $\mathbb{V}_2(S, E, I, H, V, R)$ at $(S(0), E(0), I(0), H(0), V(0), R(0))$ is

$$\mathcal{H} = \begin{pmatrix} \frac{w_5 w_1 + 1}{S^2(0)} & 0 & 0 & 0 & 0 & 0 \\ 0 & \frac{w_5 w_2}{E^2(0)} & 0 & 0 & 0 & 0 \\ 0 & 0 & \frac{w_5 w_3}{I^2(0)} & 0 & 0 & 0 \\ 0 & 0 & 0 & \frac{1}{H^2(0)} & 0 & 0 \\ 0 & 0 & 0 & 0 & \frac{w_5 w_4}{V^2(0)} & 0 \\ 0 & 0 & 0 & 0 & 0 & \frac{1}{R^2(0)} \end{pmatrix}$$

The Hessian matrix is positive definite. Thus, $\mathbb{V}_2(S, E, I, H, V, R)$ has a minimum value $\mathbb{V}_2(S(0), E(0), I(0), H(0), V(0), R(0))$.

From the continuity of $\mathbb{V}_2(S, E, I, H, V, R)$ and according to Equation (14), one can say that $\mathbb{V}_2(S, E, I, H, V, R)$ has one and only one minimum value $\mathbb{V}_2(S(0), E(0), I(0), H(0), V(0), R(0))$ in \mathbb{R}_+^6 . Now, we will define a non-negative Lyapunov \mathbb{C}^2 -function $\mathbb{V}: \mathbb{R}_+^6 \rightarrow \mathbb{R}_+$ as follows

$$\mathbb{V} = \mathbb{V}_2(S, E, I, H, V, R) - \mathbb{V}_2(S(0), E(0), I(0), H(0), V(0), R(0))$$

Applying the Itô's formula and using the stochastic model (2), we obtain

$$\begin{aligned} LV &\leq -w_5 w_6 + (w_5 w_1 + 1) \frac{\beta(E + I)(S + \sigma V)}{SN} + \lambda + 3\mu + \alpha + \gamma_3 + \rho \\ &\quad - \mu(S + E + I + H + V + R) - \sigma V - \frac{\lambda}{S} - \frac{\rho(V + R)}{S} - \frac{\epsilon I}{H} \\ &\quad - \frac{\gamma_1 E}{R} - \frac{\gamma_2 I}{R} - \frac{\gamma_3 H}{R} + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \end{aligned} \quad (16)$$

Where

$$w_6 = 5\lambda(\mathcal{R}_0^{*\frac{1}{5}} - 1) > 0$$

We next define the bounded closed set

$$\begin{aligned} D = \left\{ \varepsilon_1 < S < \frac{1}{\varepsilon_1}, \varepsilon_2 < E < \frac{1}{\varepsilon_2}, \varepsilon_3 < I < \frac{1}{\varepsilon_3}, \varepsilon_4 < H < \frac{1}{\varepsilon_4} \right. \\ \left. , \varepsilon_5 < V < \frac{1}{\varepsilon_5}, \varepsilon_6 < R < \frac{1}{\varepsilon_6} \right\} \end{aligned}$$

where $\varepsilon_i > 0, (i = 1, 2, 3, 4, 5, 6)$, we divide the whole $\mathbb{R}_+^6 \setminus D$ into the following domains

$$\begin{aligned} D_1 &= \{(S, E, I, H, V, R) \in \mathbb{R}_+^6 : 0 < S \leq \varepsilon_1\} \\ D_2 &= \{(S, E, I, H, V, R) \in \mathbb{R}_+^6 : S \geq \varepsilon_1, V \geq \varepsilon_5, 0 < I \leq \varepsilon_3\} \\ D_3 &= \{(S, E, I, H, V, R) \in \mathbb{R}_+^6 : S \geq \varepsilon_1, V < \varepsilon_5, 0 < I \leq \varepsilon_3\} \\ D_4 &= \{(S, E, I, H, V, R) \in \mathbb{R}_+^6 : S \geq \varepsilon_1, I \geq \varepsilon_3, H \geq \varepsilon_4, 0 < E < \varepsilon_2\} \end{aligned}$$

$$\begin{aligned}
D_5 &= \{(S, E, I, H, V, R) \in \mathbb{R}_+^6 : 0 < R < \varepsilon_6, I \geq \varepsilon_3, H \geq \varepsilon_4\} \\
D_6 &= \{(S, E, I, H, V, R) \in \mathbb{R}_+^6 : 0 < R < \varepsilon_6, I \geq \varepsilon_3, H < \varepsilon_4\} \\
D_7 &= \left\{ (S, E, I, H, V, R) \in \mathbb{R}_+^6 : S \geq \frac{1}{\varepsilon_1} \right\} \\
D_8 &= \left\{ (S, E, I, H, V, R) \in \mathbb{R}_+^6 : E \geq \frac{1}{\varepsilon_2} \right\} \\
D_9 &= \left\{ (S, E, I, H, V, R) \in \mathbb{R}_+^6 : I \geq \frac{1}{\varepsilon_3} \right\} \\
D_{10} &= \left\{ (S, E, I, H, V, R) \in \mathbb{R}_+^6 : C \geq \frac{1}{\varepsilon_4} \right\} \\
D_{11} &= \left\{ (S, E, I, H, V, R) \in \mathbb{R}_+^6 : V \geq \frac{1}{\varepsilon_5} \right\} \\
D_{12} &= \left\{ (S, E, I, H, V, R) \in \mathbb{R}_+^6 : R \geq \frac{1}{\varepsilon_6} \right\}
\end{aligned}$$

where $\varepsilon_i > 0$, is small enough. In the set $\mathbb{R}_+^6 \setminus D$, we can choose $\varepsilon_i > 0$ sufficiently small and satisfying

$$-\frac{\lambda}{\varepsilon_1} + C_1 \leq -j \quad (17)$$

$$-w_5 w_6 + (w_5 w_1 + 1) \frac{\beta \varepsilon_3}{N} + C_2 \leq -j \quad (18)$$

$$-w_5 w_6 + (w_5 w_1 + 1) \frac{\beta \varepsilon_3}{N} + (w_5 w_1 + 1) \frac{\beta \varepsilon_3 \sigma \varepsilon_1}{N} + C_3 \leq -j \quad (19)$$

$$(w_5 w_1 + 1) \frac{\beta \varepsilon_2}{N} + C_4 \leq -j \quad (20)$$

$$-\gamma_2 \varepsilon_6 - \gamma_3 \varepsilon_6 + C_5 \leq -j \quad (21)$$

$$-\gamma_2 \varepsilon_4 - \epsilon \varepsilon_6 + C_6 \leq -j \quad (22)$$

$$-\frac{\mu}{\varepsilon_1} + C_7 \leq -j \quad (23)$$

$$-\frac{\mu}{\varepsilon_2} + C_8 \leq -j \quad (24)$$

$$-\frac{\mu}{\varepsilon_3} + C_9 \leq -j \quad (25)$$

$$-\frac{\mu}{\varepsilon_4} + C_{10} \leq -j \quad (26)$$

$$-\frac{(\mu + \sigma)}{\varepsilon_5} + C_{11} \leq -j \quad (27)$$

$$-\frac{\mu}{\varepsilon_6} + C_{12} \leq -j \quad (28)$$

Next, we will show that $L\mathbb{V}(S, E, I, H, V, R) \leq -j$ on $\mathbb{R}_+^6 \setminus D$, which is equivalent to proving it on the above twelfth domains.

Case 1. If $(S, E, I, H, V, R) \in D_1$, then by equation (16) we get

$$\begin{aligned} L\mathbb{V} &\leq -w_5w_6 + (w_5w_1 + 1)\frac{\beta(E+I)(S+\sigma V)}{SN} + \lambda + 3\mu + \alpha + \gamma_3 + \rho \\ &\quad - \mu(S + E + I + H + V + R) - \sigma V - \frac{\lambda}{S} - \frac{\rho(V+R)}{S} - \frac{\epsilon I}{H} \\ &\quad - \frac{\gamma_1 E}{R} - \frac{\gamma_2 I}{R} - \frac{\gamma_3 H}{R} + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \\ &\leq -\frac{\lambda}{S} + C_1 \leq -\frac{\lambda}{\varepsilon_1} + C_1 \end{aligned}$$

where

$$\begin{aligned} C_1 = \sup_{(S,E,I,H,V,R) \in \mathbb{R}_+^6} &\left\{ -w_5w_6 + (w_5w_1 + 1)\frac{\beta(E+I)(S+\sigma V)}{SN} + \lambda + 3\mu \right. \\ &\quad \left. + \alpha + \gamma_3 + \rho - \mu(S + E + I + H + V + R) - \sigma V - \frac{\rho(V+R)}{S} \right. \\ &\quad \left. - \frac{\epsilon I}{H} - \frac{\gamma_1 E}{R} - \frac{\gamma_2 I}{R} - \frac{\gamma_3 H}{R} + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \right\} \end{aligned}$$

According to (17), we have $L\mathbb{V} \leq -j$ for any $(S, E, I, H, V, R) \in D_1$.

Case 2. If $(S, E, I, H, V, R) \in D_2$, then by equation (16) we get

$$L\mathbb{V} \leq -w_5w_6 + (w_5w_1 + 1)\frac{\beta I}{N} + C_2 \leq -w_5w_6 + (w_5w_1 + 1)\frac{\beta \varepsilon_3}{N} + C_2$$

where

$$\begin{aligned} C_2 = \sup_{(S,E,I,H,V,R) \in \mathbb{R}_+^6} &\left\{ (w_5w_1 + 1)\frac{\beta E(S+\sigma V)}{SN} + (w_5w_1 + 1)\frac{\beta I\sigma V}{SN} + \lambda \right. \\ &\quad \left. + 3\mu + \alpha + \gamma_3 + \rho - \mu(S + E + I + H + V + R) - \sigma V - \frac{\lambda}{S} \right. \\ &\quad \left. - \frac{\rho(V+R)}{S} - \frac{\epsilon I}{H} - \frac{\gamma_1 E}{R} - \frac{\gamma_2 I}{R} - \frac{\gamma_3 H}{R} + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \right\} \end{aligned}$$

According to (18), we have $L\mathbb{V} \leq -j$ for any $(S, E, I, H, V, R) \in D_2$.

Case 3. If $(S, E, I, H, V, R) \in D_3$, then

$$\begin{aligned} L\mathbb{V} &\leq -w_5w_6 + (w_5w_1 + 1)\frac{\beta I}{N} + (w_5w_1 + 1)\frac{\beta I\sigma V}{SN} + C_3 \\ &\leq -w_5w_6 + (w_5w_1 + 1)\frac{\beta \varepsilon_3}{N} + (w_5w_1 + 1)\frac{\beta \varepsilon_3 \sigma \varepsilon_5}{\varepsilon_1 N} + C_3 \\ &\leq -w_5w_6 + (w_5w_1 + 1)\frac{\beta \varepsilon_3}{N} + (w_5w_1 + 1)\frac{\beta \varepsilon_3 \sigma \varepsilon_1}{N} + C_3 \end{aligned}$$

Choosing $\varepsilon_5 = \varepsilon_1^2$, where

$$C_3 = \sup_{(S,E,I,H,V,R) \in \mathbb{R}_+^6} \left\{ (w_5w_1 + 1)\frac{\beta E(S+\sigma V)}{SN} + \lambda + 3\mu + \alpha + \gamma_3 + \rho \right.$$

$$-\mu(S + E + I + H + V + R) - \sigma V - \frac{\lambda}{S} - \frac{\rho(V + R)}{S} - \frac{\epsilon I}{H} - \frac{\gamma_1 E}{R} - \frac{\gamma_2 I}{R} - \frac{\gamma_3 H}{R} + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \Big\}$$

According to (19), we have $L\mathbb{V} \leq -j$ for any $(S, E, I, H, V, R) \in D_3$.

Case 4. If $(S, E, I, H, V, R) \in D_4$, then

$$L\mathbb{V} \leq (w_5 w_1 + 1) \frac{\beta E}{N} + C_4 \leq (w_5 w_1 + 1) \frac{\beta \varepsilon_2}{N} + C_4$$

where

$$\begin{aligned} C_4 = & \sup_{(S, E, I, H, V, R) \in \mathbb{R}_+^6} \left\{ -w_5 w_6 + (w_5 w_1 + 1) \frac{\beta I(S + \sigma V)}{SN} + \lambda + 3\mu + \alpha \right. \\ & + \gamma_3 + \rho + (w_5 w_1 + 1) \frac{\beta E \sigma V}{SN} - \mu(S + E + I + H + V + R) \\ & - \sigma V - \frac{\lambda}{S} - \frac{\rho(V + R)}{S} - \frac{\epsilon I}{H} - \frac{\gamma_1 E}{R} - \frac{\gamma_2 I}{R} - \frac{\gamma_3 H}{R} \\ & \left. + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \right\} \end{aligned}$$

By (20), we can conclude that $L\mathbb{V} \leq -j$ on D_4 .

Case 5. If $(S, E, I, H, V, R) \in D_5$, then

$$\begin{aligned} L\mathbb{V} & \leq -\frac{\gamma_2 I}{R} - \frac{\gamma_3 H}{R} + C_5 \\ & \leq -\frac{\gamma_2 \varepsilon_3}{\varepsilon_6} - \frac{\gamma_3 \varepsilon_4}{\varepsilon_6} + C_5 \\ & \leq -\gamma_2 \varepsilon_6 - \gamma_3 \varepsilon_6 + C_5 \end{aligned}$$

Choosing $\varepsilon_3 = \varepsilon_6^2$, $\varepsilon_4 = \varepsilon_6^2$

where

$$\begin{aligned} C_5 = & \sup_{(S, E, I, H, V, R) \in \mathbb{R}_+^6} \left\{ -w_5 w_6 + (w_5 w_1 + 1) \frac{\beta(E + I)(S + \sigma V)}{SN} + \lambda + 3\mu \right. \\ & + \alpha + \gamma_3 + \rho - \mu(S + E + I + H + V + R) - \sigma V - \frac{\lambda}{S} \\ & \left. - \frac{\rho(V + R)}{S} - \frac{\epsilon I}{H} - \frac{\gamma_1 E}{R} + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \right\} \end{aligned}$$

By (21), we can conclude that $L\mathbb{V} \leq -j$ on D_5 .

Case 6. If $(S, E, I, H, V, R) \in D_6$, then

$$L\mathbb{V} \leq -\frac{\gamma_2 I}{R} - \frac{\epsilon I}{H} + C_6$$

$$\begin{aligned} &\leq -\frac{\gamma_2 \varepsilon_3}{\varepsilon_6} - \frac{\varepsilon \varepsilon_3}{\varepsilon_4} + C_6 \\ &\leq -\gamma_2 \varepsilon_4 - \varepsilon \varepsilon_6 + C_6 \end{aligned}$$

Choosing $\varepsilon_3 = \varepsilon_6 \varepsilon_4$

where

$$\begin{aligned} C_6 = & \sup_{(S,E,I,H,V,R) \in \mathbb{R}_+^6} \left\{ -w_5 w_6 + (w_5 w_1 + 1) \frac{\beta(E+I)(S+\sigma V)}{SN} + \lambda + 3\mu \right. \\ & + \alpha + \gamma_3 + \rho - \mu(S+E+I+H+V+R) - \sigma V - \frac{\lambda}{S} \\ & \left. - \frac{\rho(V+R)}{S} - \frac{\gamma_3 H}{R} - \frac{\gamma_1 E}{R} + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \right\} \end{aligned}$$

According to (22), we have $L\mathbb{V} \leq -j$ for any $(S, E, I, H, V, R) \in D_6$.

Case 7. If $(S, E, I, H, V, R) \in D_7$, then

$$\begin{aligned} L\mathbb{V} &\leq -w_5 w_6 + (w_5 w_1 + 1) \frac{\beta(E+I)(S+\sigma V)}{SN} + \lambda + 3\mu + \alpha + \gamma_3 \\ &\quad + \rho - \mu(S+E+I+H+V+R) - \sigma V - \frac{\lambda}{S} - \frac{\rho(V+R)}{S} \\ &\quad - \frac{\varepsilon I}{H} - \frac{\gamma_1 E}{R} - \frac{\gamma_2 I}{R} - \frac{\gamma_3 H}{R} + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \\ &\leq -\mu S + C_7 \leq -\frac{\mu}{\varepsilon_1} + C_7 \end{aligned}$$

where

$$\begin{aligned} C_7 = & \sup_{(S,E,I,H,V,R) \in \mathbb{R}_+^6} \left\{ -w_5 w_6 + (w_5 w_1 + 1) \frac{\beta(E+I)(S+\sigma V)}{SN} + \lambda \right. \\ & + 3\mu + \alpha + \gamma_3 + \rho - \mu(E+I+H+V+R) - \sigma V - \frac{\lambda}{S} \\ & \left. - \frac{\rho(V+R)}{S} - \frac{\varepsilon I}{H} - \frac{\gamma_1 E}{R} - \frac{\gamma_2 I}{R} - \frac{\gamma_3 H}{R} + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \right\} \end{aligned}$$

According to (23), we have $L\mathbb{V} \leq -j$ for any $(S, E, I, H, V, R) \in D_7$.

Case 8. If $(S, E, I, H, V, R) \in D_8$, then

$$L\mathbb{V} \leq -\mu E + C_8 \leq -\frac{\mu}{\varepsilon_2} + C_8$$

where

$$\begin{aligned} C_8 = & \sup_{(S,E,I,H,V,R) \in \mathbb{R}_+^6} \left\{ -w_5 w_6 + (w_5 w_1 + 1) \frac{\beta(E+I)(S+\sigma V)}{SN} + \lambda \right. \\ & \left. + 3\mu + \alpha + \gamma_3 + \rho - \mu(S+I+H+V+R) - \sigma V - \frac{\lambda}{S} \right\} \end{aligned}$$

$$-\frac{\rho(V+R)}{S} - \frac{\epsilon I}{H} - \frac{\gamma_1 E}{R} - \frac{\gamma_2 I}{R} - \frac{\gamma_3 H}{R} + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \Bigg\}$$

According to (24), we have $L\mathbb{V} \leq -j$ for any $(S, E, I, H, V, R) \in D_8$.

Case 9. If $(S, E, I, H, V, R) \in D_9$, then

$$L\mathbb{V} \leq -\mu I + C_9 \leq -\frac{\mu}{\varepsilon_3} + C_9$$

where

$$\begin{aligned} C_9 = & \sup_{(S,E,I,H,V,R) \in \mathbb{R}_+^6} \left\{ -w_5 w_6 + (w_5 w_1 + 1) \frac{\beta(E+I)(S+\sigma V)}{SN} + \lambda \right. \\ & + 3\mu + \alpha + \gamma_3 + \rho - \mu(S+E+H+V+R) - \sigma V - \frac{\lambda}{S} \\ & \left. - \frac{\rho(V+R)}{S} - \frac{\epsilon I}{H} - \frac{\gamma_1 E}{R} - \frac{\gamma_2 I}{R} - \frac{\gamma_3 H}{R} + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \right\} \end{aligned}$$

According to (25), we have $L\mathbb{V} \leq -j$ for any $(S, E, I, H, V, R) \in D_9$.

Case 10. If $(S, E, I, H, V, R) \in D_{10}$, then

$$L\mathbb{V} \leq -\mu H + C_{10} \leq -\frac{\mu}{\varepsilon_4} + C_{10}$$

where

$$\begin{aligned} C_{10} = & \sup_{(S,E,I,H,V,R) \in \mathbb{R}_+^6} \left\{ -w_5 w_6 + (w_5 w_1 + 1) \frac{\beta(E+I)(S+\sigma V)}{SN} + \lambda \right. \\ & + 3\mu + \alpha + \gamma_3 + \rho - \mu(S+E+I+V+R) - \sigma V - \frac{\lambda}{S} \\ & \left. - \frac{\rho(V+R)}{S} - \frac{\epsilon I}{H} - \frac{\gamma_1 E}{R} - \frac{\gamma_2 I}{R} - \frac{\gamma_3 H}{R} + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \right\} \end{aligned}$$

According to (26), we have $L\mathbb{V} \leq -j$ for any $(S, E, I, H, V, R) \in D_{10}$.

Case 11. If $(S, E, I, H, V, R) \in D_{11}$, then

$$L\mathbb{V} \leq -\mu V - \sigma V + C_{11} \leq -\frac{(\mu + \sigma)}{\varepsilon_5} + C_{11}$$

where

$$\begin{aligned} C_{11} = & \sup_{(S,E,I,H,V,R) \in \mathbb{R}_+^6} \left\{ -w_5 w_6 + (w_5 w_1 + 1) \frac{\beta(E+I)(S+\sigma V)}{SN} + \lambda \right. \\ & + 3\mu + \alpha + \gamma_3 + \rho - \mu(S+E+I+H+R) - \frac{\rho(V+R)}{S} \\ & \left. - \frac{\lambda}{S} - \frac{\epsilon I}{H} - \frac{\gamma_1 E}{R} - \frac{\gamma_2 I}{R} - \frac{\gamma_3 H}{R} + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \right\} \end{aligned}$$

According to (27), we have $L\mathbb{V} \leq -j$ for any $(S, E, I, H, V, R) \in D_{11}$.

Case 12. If $(S, E, I, H, V, R) \in D_{12}$, then

$$L\mathbb{V} \leq -\mu R + C_{12} \leq -\frac{\mu}{\varepsilon_6} + C_{12}$$

where

$$C_{12} = \sup_{(S, E, I, H, V, R) \in \mathbb{R}_+^6} \left\{ -w_5 w_6 + (w_5 w_1 + 1) \frac{\beta(E + I)(S + \sigma V)}{SN} + \lambda \right. \\ \left. + 3\mu + \alpha + \gamma_3 + \rho - \mu(S + E + I + H + V) - \sigma V - \frac{\lambda}{S} \right. \\ \left. - \frac{\rho(V + R)}{S} - \frac{\epsilon I}{H} - \frac{\gamma_1 E}{R} - \frac{\gamma_2 I}{R} - \frac{\gamma_3 H}{R} + \frac{\delta_1^2 + \delta_4^2 + \delta_6^2}{2} \right\}$$

According to (28), we have $L\mathbb{V} \leq -j$ for any $(S, E, I, H, V, R) \in D_{12}$. Obviously, from equations (17)–(28), one can see for a sufficiently small ε_i that

$$L\mathbb{V}(S, E, I, H, V, R) \leq -j \quad \text{for all } (S, E, I, H, V, R) \in \mathbb{R}_+^6 \setminus D.$$

Consequently, condition two in Lemma 4.1 in [41] is satisfied. This show that stochastic model (2) is ergodic and has a unique stationary distribution.

4. Numerical Simulations

This section concerns the estimation of parameters for the deterministic model (1) through weekly data of reported seasonal influenza cases from Saudi Arabia in 2022. The data were obtained from the official WHO database for influenza surveillance[42]. Parameters have been estimated via a Bayesian inference approach and the Monte Carlo Markov chain methodology [43–45]. This facilitates the assessment of parameter uncertainty while calibrating the model to the observed time series data.

Table 2: Estimated values from the data Influenza

parameters	Estimated value	Reference
λ	100	Fitted
μ	0.01246293	Fitted
β	0.36983069	Fitted
α	0.18302843	Fitted
σ	0.89689644	Fitted
ϵ	0.69978458	Fitted
γ_1	0.16348621	Fitted
γ_2	0.49999414	Fitted
γ_3	0.18814877	Fitted
ω	0.10000064	Fitted
ρ	0.00038610	Fitted

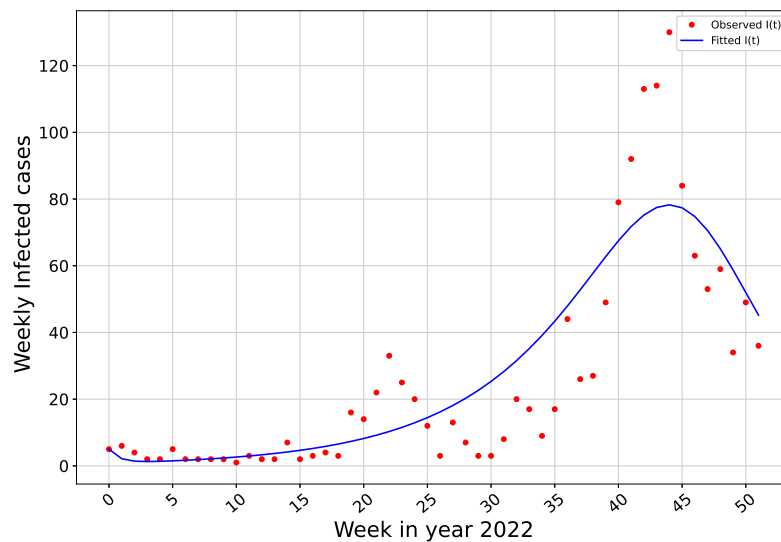


Figure 1: Model fit to weekly influenza cases in Saudi Arabia, 2022.

The fit of the deterministic model model(1) to the weekly reported influenza cases in Saudi Arabia for the year 2022 is shown in Figure1. The observed data spots are marked by red dots, and the model trend is drawn as the blue curve. The. But it does capture very well the epidemic nature. The seasonal peak (height during epidemiological week 30) appears well and is well described by the model. This alignment emphasizes the model's potential for identifying the grand epidemiological profile, as well as for yielding guidance for public health interventions.

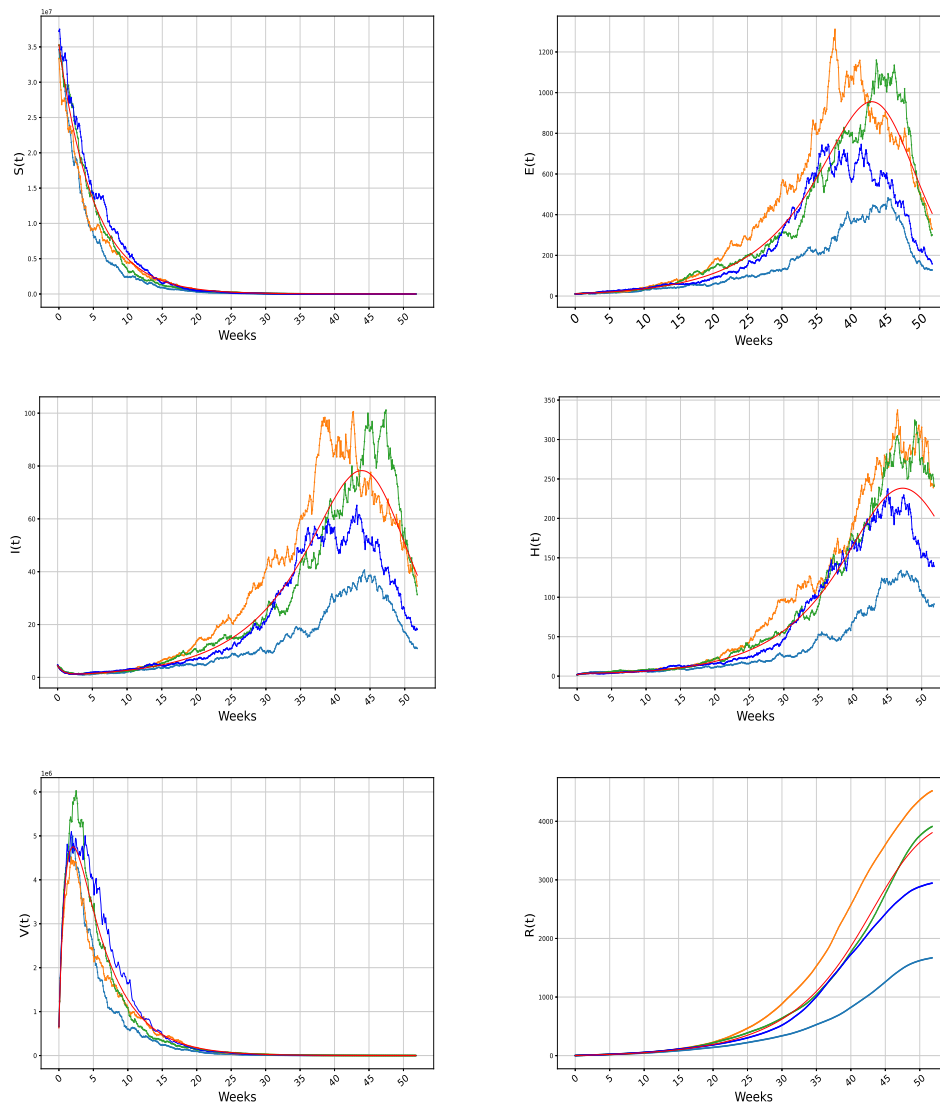


Figure 2: Stochastic simulations vs. deterministic model (1) (red line) using estimated parameters from Table 2 with $\delta_i = 0.1$.

Figure 2 shows how the determined and the stochastic model results are associated in six compartments of the influenza system in 52 weeks. The deterministic trajectories are represented as red solid lines, indicating a singular solution trajectory derived from the average behaviour. The stochastic encoding (for $\delta_1 = \delta_2 = \delta_3 = \delta_4 = \delta_5 = \delta_6 = 0.1$), on the other hand, provides a distribution of potential epidemic outcomes, applying randomness coming from fluctuations in the environment and the population behavior. The strong alignment between the outcomes of the stochastic framework and the deterministic solution indicates that the mean-field model is valid in the low-noise limit. The findings indicate

that the stochastic model effectively encapsulates the uncertainty and variability inherent in the epidemic's progression, offering a more comprehensive understanding of the potential dynamics of the outbreak.

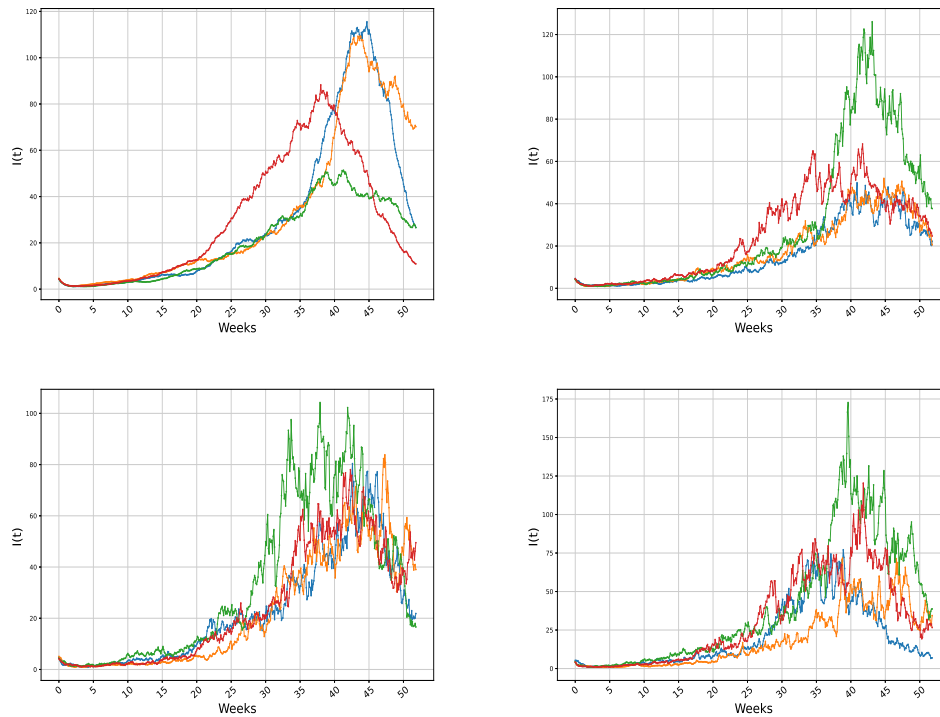


Figure 3: Stochastic simulations of model (2) with $\delta_1 = \delta_2 = \delta_5 = \delta_6 = 0.1$ and varying $\delta_3 = \delta_4 \in \{0.05, 0.2, 0.25, 0.3\}$.

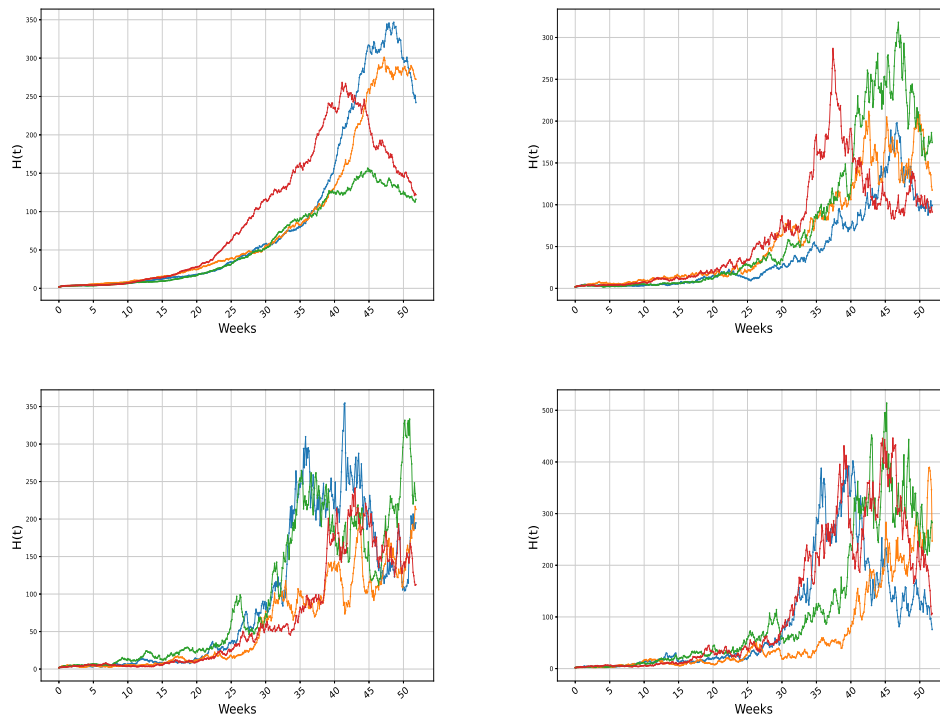


Figure 4: Stochastic simulations of model (2) with $\delta_1 = \delta_2 = \delta_5 = \delta_6 = 0.1$ and varying $\delta_3 = \delta_4$: 0.05, 0.2 (top), 0.25, 0.3 (bottom).

The effect of varying levels of white noise intensities (δ_3, δ_4) on the dynamics of influenza transmission over 52 weeks is presented using stochastic simulations in Figures 3 and 4. Each subplot shows simulated trajectories of the epidemic for a range of noise levels in panel (a) and panel (i) with the x-axis the time in weeks and the y-axis the infected or hospitalized counts. We set $\delta_1 = 0.1$, $\delta_2 = 0.1$, $\delta_5 = 0.1$ and $\delta_6 = 0.1$ and for the sensitivity analysis we vary δ_3 and δ_4 systematically.

With increasing noise intensity ($0.05 \leq \delta_3 = \delta_4 \leq 0.3$), the variation in trajectories also increases, indicating increasing prediction uncertainty for both infection and hospitalization dynamics. This diffusion shows a method's dependency on randomly occurring perturbations, highlighting the essential significance of randomness originating from the environment and behavior. In particular, larger levels of noise result in more deviation from the deterministic mean curve, showing that randomness may intensify or dampen the impact of outbreaks depending on the initial condition and the system behavior. These results underscore the need to consider uncertainty when developing epidemic forecasting models and devising resilient public health responses.

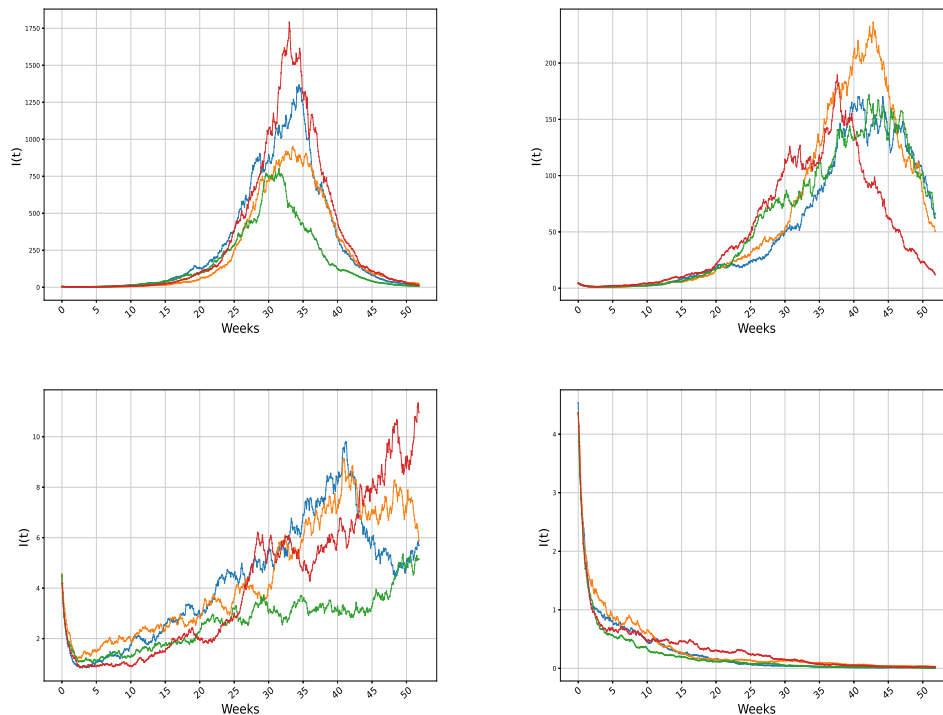


Figure 5: Stochastic simulations with varying β : Top— $\beta = 0.5$ and 0.4 yield higher peaks (801–1790 and 170–236 cases); Bottom— $\beta = 0.3$ shows minimal infections, while $\beta = 0.2$ leads to extinction.

Figures 5 & 6 show the effect of decreasing the transmission rate β on infection and hospitalization dynamics over 52 weeks of simulation.

Figure 5 shows the trajectories of the infection:

- With $\beta = 0.5$, we observed infections peak during weeks 32–35 with between 801 to 1790 cases.
- When $\beta = 0.4$, the peak is shifted to between weeks 38–43, with a height between 170–236 cases.
- Infectious peaks are much lower and estimated at 11 cases at $\beta = 0.3$.
- When $\beta = 0.2$, the infection dies out: it goes extinct.

Figure 6 shows the same statistics for hospitalizations:

- The parameter values that provide $\beta = 0.5$ and $\beta = 0.4$ lead to the cases peaking from 2305 to 4659 in hospitals.
- At $\beta = 0.3$, there are at most 32 new hospitalisations in a week.

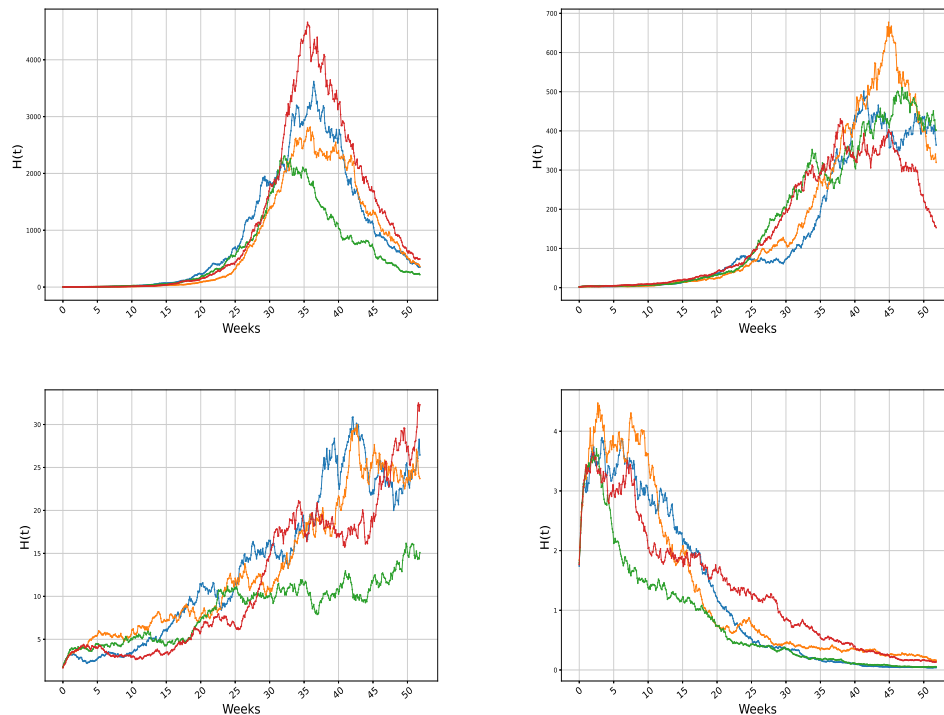


Figure 6: Stochastic simulations of model (2) for varying transmission rates β . Top: (left) $\beta = 0.5$, hospitalizations peak between 2305 and 4659 cases; (right) $\beta = 0.4$, similar peak range observed. Bottom: (left) $\beta = 0.3$, peak hospital cases limited to 32; (right) $\beta = 0.2$, the epidemic fades out.

- For $\beta = 0.2$ the disease vanishes completely.

These simulations verify that large β causes early and strong outbreaks, while small β suppresses transmission. This sensitivity analysis highlights the importance of the transmission parameter to determine outbreak size and guide intervention efforts under the public health strategy.

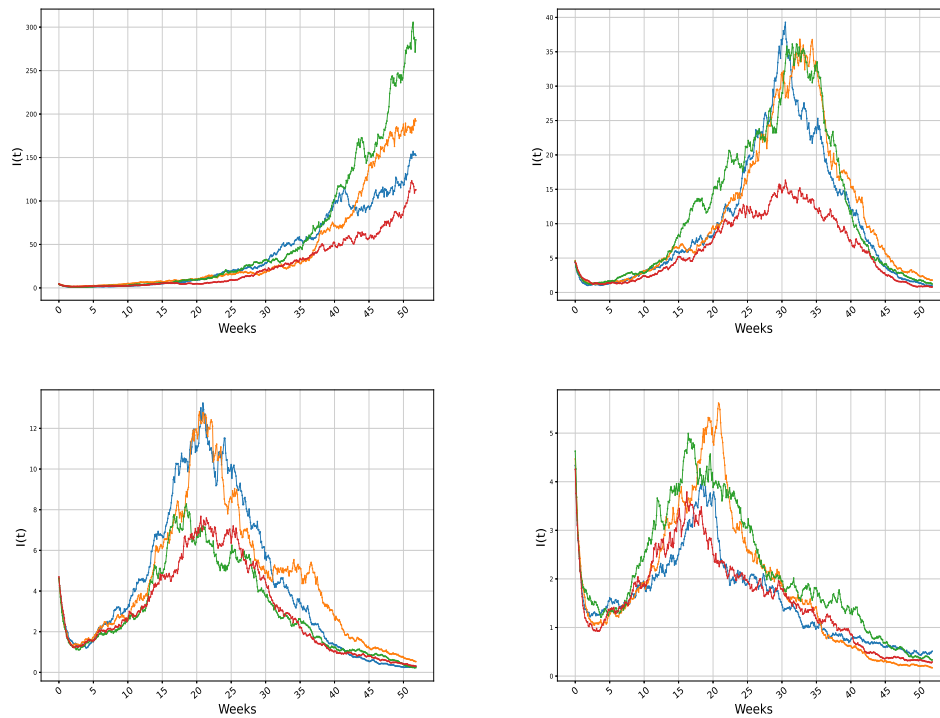


Figure 7: Stochastic simulations of model (2) for varying α . As α increases from 0.1 to 0.7, peak infections occur earlier and total cases decrease, indicating faster disease extinction.

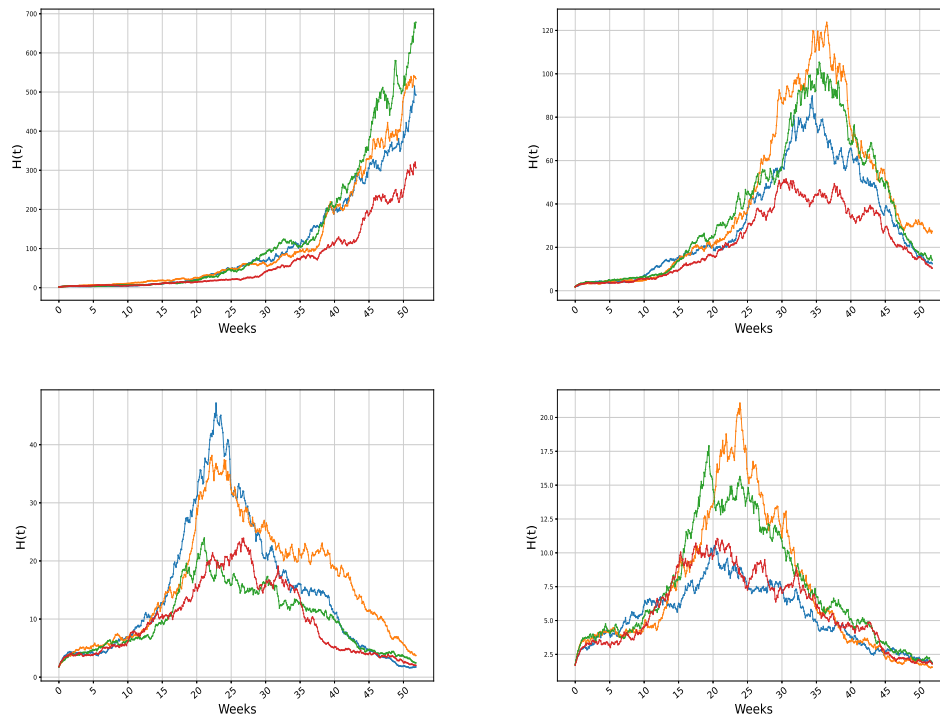


Figure 8: Stochastic model (2) with varying α . Higher α leads to earlier, smaller peaks in hospital cases, decreasing from 678 ($\alpha = 0.1$) to 21 ($\alpha = 0.7$).

The effect of the recovery rate α is depicted in **Figures 7–10** on the infection and hospitalization dynamics for 52 weeks.

Infection paths as α increases are visualized in **Figure 7**. For $\alpha = 0.1$, infections are maintained with number of the cases ranging from 123 to 305. In this case, for $\alpha = 0.5, 0.7$, the peak moves to an earlier time and the total infection decreases albeit naturally (the recovery occurs faster).

Hospitalizations are presented in Fig 8. For $\alpha = 0.1$, the peak hospitalizations range from 320 to 678. Larger α values give rise to lower and earlier peaks, while deaths in hospital have decreased considerably by $\alpha = 0.7$.

9 dives into infection heterogeneities at $\alpha = 0.9, 0.7, 0.5$, and 0.3 . A larger α corresponds to sharper peaks and peaks that occur earlier in time, and also to fewer total cases, emphasizing the role of rapid recovery in outbreak suppression.

Figure 10 verifies the above pattern for hospital cases. With $\alpha = 0.3$ and 0.5 , hospital burdens are small, but $\alpha = 0.9$ leads to brief yet intense peaks.

Collectively, higher α leads to a lower epidemic duration and severity, which emphasizes it as a key factor that affects both infection and hospitalization outcomes. These findings confirm the relevance of the stochastic model in representing the time and magnitude of epidemic fadeout at different recovery settings.

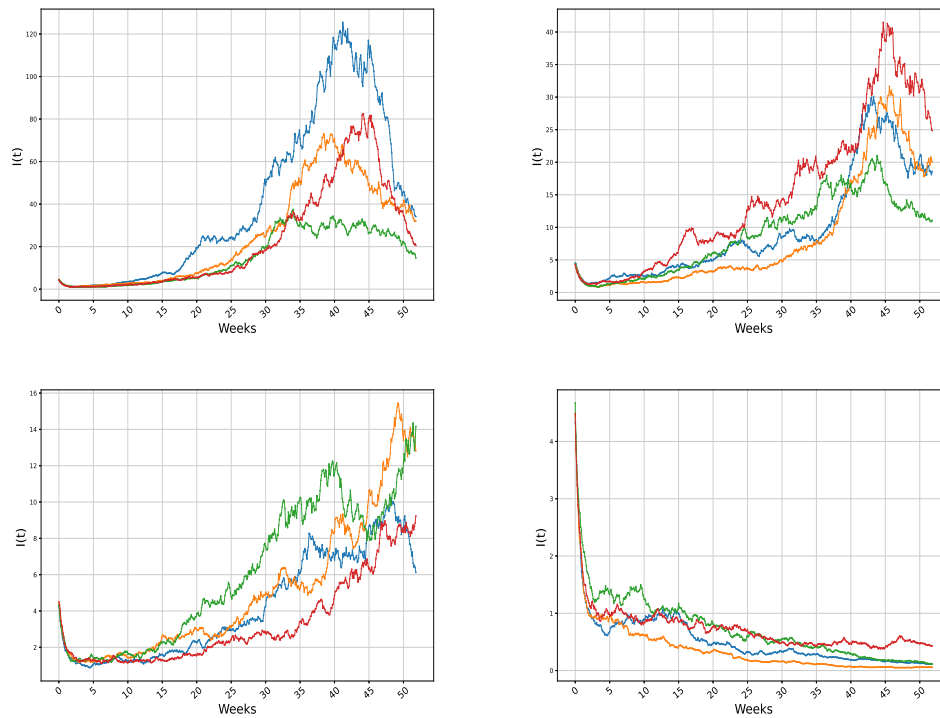


Figure 9: Stochastic model (2) for different σ values. As σ decreases from 0.9 to 0.3, infection peaks occur later and at lower levels, with extinction observed at $\sigma = 0.3$.

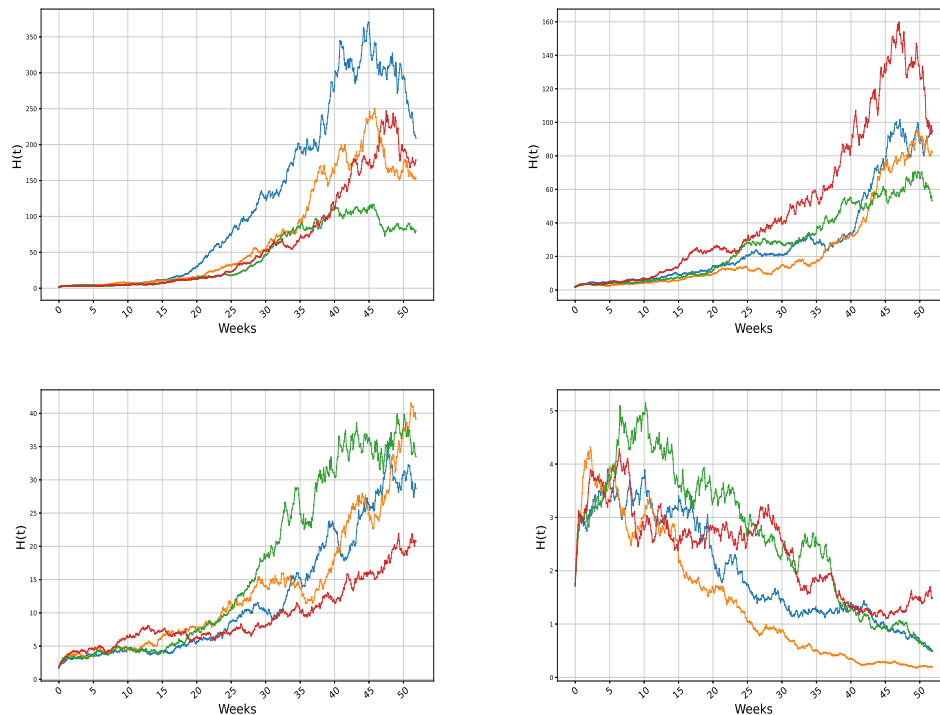


Figure 10: Stochastic model (2) with varying σ . As σ decreases from 0.9 to 0.3, hospital cases peak later and decline in magnitude, with extinction at $\sigma = 0.3$.

5. Conclusion

This study uses a stochastic model to assess Saudi Arabia's weekly seasonal influenza cases in 2022. The model closely aligns with the general trend of observed influenza cases, showing a peak in infections around week 30, which suggests a seasonal epidemic. When comparing deterministic and stochastic models, it becomes evident that the stochastic model excels at capturing a broader spectrum of potential outcomes due to its natural variability. This study delves into the substantial influence of white noise intensity on the variability of infection predictions within the stochastic model, highlighting the crucial role of stochastic factors like environmental or social influences in modeling endeavors. Moreover, the parameter β plays a crucial role in impacting infection and hospitalization rates, as higher β values are linked to more severe epidemic outcomes. Finally, the parameter α plays a vital role in forecasting the timing and intensity of infection and hospitalization peaks. Higher α values result in an earlier peak and a swift decrease in cases, which may be associated with recovery rates or other factors that help resolve the disease. These findings demonstrate the stochastic model's sensitivity to critical factors needed to predict illness progression and help build effective public health treatments and epidemic control measures. In the future, we intend to solve some new models, such as in [46–48] and make comparisons with other methods [49–53].

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Conflicts of Interest: The authors declare that they have no conflict of interest.

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